

1 Quantifying the known unknowns: estimating
2 maximum intrinsic rate of population increase in the
3 face of uncertainty

4 Sebastián A. Pardo*

5 Department of Biology, Dalhousie University, Halifax, NS, B3H 4R2,
6 Canada

7 Earth to Ocean Research Group, Biological Sciences, Simon Fraser
8 University, Burnaby, BC, V5A 1S6, Canada

9
10 Andrew B. Cooper

11 School of Resource & Environmental Management, Simon Fraser
12 University, Burnaby, BC, V5A 1S6, Canada

13
14 John D. Reynolds

15 Earth to Ocean Research Group, Biological Sciences, Simon Fraser
16 University, Burnaby, BC, V5A 1S6, Canada

17
18 Nicholas K. Dulvy

19 Earth to Ocean Research Group, Biological Sciences, Simon Fraser
20 University, Burnaby, BC, V5A 1S6, Canada

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23 **Abstract**

24 Scientists are often called upon to provide advice on the consequences of exploiting
25 data-poor and potentially threatened sharks and rays. For species lacking detailed
26 biological information, sensitivity to overfishing is often estimated using simple
27 models that depend upon life history parameters. Yet, there has been little explo-

28 ration of the degree to which uncertainty in life history parameters can influence
29 demographic parameter estimates and therefore fisheries management options. We
30 estimate the maximum intrinsic rate of population increase (r_{max}) for ten coastal
31 carcharhiniform shark populations using an unstructured life history model that
32 explicitly accounts for uncertainty in life history parameters. We also seek to un-
33 derstand which parameters most influenced the estimate of r_{max} . The median r_{max}
34 estimate proved to be robust to the propagation of uncertainty of life history param-
35 eters when that uncertainty is symmetric around the parameters' point estimates,
36 but accounting for uncertainty can lead to a wide range of plausible r_{max} estimates
37 for any given species. While natural mortality M can considerably impact r_{max}
38 estimates when its uncertainty is high it is very difficult to estimate this param-
39 eter directly, hence we focused on the two directly estimated parameters: age at
40 maturity α_{mat} and annual reproductive output b . Uncertainty in age at maturity
41 values was low, yet it still resulted in moderate uncertainty in r_{max} estimates. The
42 model was particularly sensitive to uncertainty in annual reproductive output as
43 b approached values of less than 5 female offspring per year, which is not unusual
44 for large elasmobranchs and marine mammals. Furthermore, at very low b values
45 (<1), there is a threshold that results in implausible r_{max} estimates when both
46 M and α_{mat} are high. Managers and policy makers should be careful to restrict
47 mortality on species with very low annual reproductive output <2 females per year.
48 We recommend elasmobranch biologists to measure frequency distributions of litter
49 sizes (rather than just a range) and age at maturity, as well as improving estimates
50 of natural mortality of data-poor elasmobranchs.

51 **Keywords:** Bycatch, Carcharhinus, Chondrichthyes, Elasmobranchii, reference points, risk
52 assessment, demography

53 1 Introduction

54 Many marine megafauna and predator populations are declining globally and are at in-
55 creasing risk of local and regional extinction (Fowler *et al.*, 2005; Christensen *et al.*, 2014;
56 Dulvy *et al.*, 2017). Sharks, rays, and chimaeras (class Chondrichthyes) play a complex
57 role as marine predators, hence there is concern for potential indirect impacts of predator
58 declines on marine ecosystems (Kitchell *et al.*, 2002; Heithaus *et al.*, 2008; Heupel *et al.*,
59 2014). They are often large-bodied, long-lived, late-maturing and produce few offspring
60 (Compagno, 1990; Musick, 1999). Consequently, they tend to have low intrinsic rates
61 of population increase and weak compensatory density dependence in juvenile survival
62 (Compagno, 1990; Musick, 1999; Forrest and Walters, 2009); traits which make them less
63 able to withstand over-exploitation (Kindsvater *et al.*, 2016).

*Corresponding author: email: spardo@dal.ca

64 There are relatively few stock-assessed elasmobranch species (Simpfendorfer and Dulvy,
65 2017) even though they can be directly targeted (Jaiteh *et al.*, 2016, 2017) and are also
66 caught as bycatch in longline fisheries targeting large finfish, coastal gill-net, and trawl
67 fisheries (Kitchell *et al.*, 2002; Ellis *et al.*, 2005). Despite the large catch and high commer-
68 cial value of elasmobranchs, our understanding of species-specific catches is poor because
69 many species are not directly targeted (Stevens, 2000) but are considered valuable by-
70 catch. Hence, accurate information on population trends is lacking for most species.
71 These challenges to understanding sustainability of elasmobranch fisheries and using pre-
72 cautionary approaches for their management are compounded further by multi-species
73 fisheries and poor species-specific monitoring (Barker and Schluessel, 2005; Lack and
74 Sant, 2009; Dulvy *et al.*, 2017).

75 Many countries have recently adopted policy regulations that require them to assess
76 fisheries according to an Ecosystem-Based Management (EBM) approach (Jennings and
77 Rice, 2011; Rogers *et al.*, 2007). The EBM approach requires, among other things, the
78 identification of safe ecological limits for bycatch species (Hobday *et al.*, 2011; Salomon
79 and Holm-Müller, 2013). These species, which lack detailed stock assessments and are
80 not the focus of targeted commercial extraction, are usually understudied, resulting in a
81 dearth of information on their biology and demography. As such, the usual data-intensive
82 stock assessment methods are not applicable for a large diversity of bycatch, which has
83 led to a recent increase in the development of tools for the assessment of data-poor species
84 (Brooks *et al.*, 2010).

85 Identifying which life history traits affect resilience to a range of fishing pressures is
86 crucial for averting over-exploitation or extinction of data-poor species (Reynolds, 2003;
87 Kindsvater *et al.*, 2016). Life history traits are interrelated due to the evolutionary con-
88 straints imposed by energy acquisition and processing (Law, 1979; Charnov, 1993). Some
89 of these relationships, widely known as Beverton-Holt dimensionless ratios, can be used
90 to predict other life history parameters and tied to population dynamics, albeit with
91 considerable uncertainty (Dulvy and Forrest, 2010). The link between life histories and
92 demography allows the use of life history traits to quantify a species' intrinsic sensitivity
93 (Frisk *et al.*, 2001; Dulvy *et al.*, 2004; Reynolds *et al.*, 2005), which encompasses biological
94 traits that, in conjunction with exposure to threatening processes, determine their vul-
95 nerability or extinction risk (Reynolds, 2003; Dulvy and Kindsvater, 2017). For example,
96 age- and stage-structured models have been used to estimate the intrinsic sensitivity of
97 numerous shark and ray species (e.g. Cortés, 2002; Mollet and Cailliet, 2002; Simpfendor-
98 fer, 2005). Such models depend heavily on age- and stage-specific estimates of growth,
99 natural mortality and reproductive output, but such detailed information is often lack-
100 ing for most elasmobranchs, particularly natural mortality (Miller *et al.*, 2003; Gedamke
101 *et al.*, 2007). Alternatively, unstructured models do not require age- or stage-specific life
102 history estimates and instead use single trait estimates. Unstructured models have the

103 advantage of relying on simple assumptions on how fertility or survival may vary with
104 age or stage, and have been used to estimate relative intrinsic sensitivity for numerous
105 data-poor elasmobranch species (Smith *et al.*, 1998; García *et al.*, 2008; Simpfendorfer
106 and Kyne, 2009).

107 Recently, there has been increasing awareness of the importance of considering mul-
108 tiple sources of uncertainty in demographic parameter estimation and risk assessment
109 (Simpfendorfer *et al.*, 2011; Cortés *et al.*, 2014; Jaiteh *et al.*, 2016). In addition, demo-
110 graphic modelling frameworks quantify the degree of caution that should be exercised for
111 their sustainable management and can have major implications for the conservation of
112 species (Caswell *et al.*, 1998; Cortés, 2002; Cortés *et al.*, 2014). The two main sources
113 of uncertainty that can be easily accounted for in a modelling framework are measure-
114 ment error (or trait error), stemming from uncertainty in the empirical estimation of a
115 life history parameter (Harwood and Stokes, 2003; Quiroz *et al.*, 2010), and coefficient
116 error, which is derived from the uncertainty in the values of the coefficients of a model
117 (Quiroz *et al.*, 2010, e.g., uncertainty around the intercept of a linear model, see). While
118 multiple sources of uncertainty can be readily accounted for in stock assessments, this
119 has not happened to the same extent in data-poor situations, particularly in commonly
120 used unstructured models (for a recent example see Jaiteh *et al.*, 2016).

121 In this study, we use an unstructured derivation of the Euler-Lotka demographic
122 model, which estimates the maximum intrinsic rate of population increase r_{max} (Myers
123 *et al.*, 1997; Pardo *et al.*, 2016; Cortés, 2016). We address how measurement error in
124 life history traits affects (1) uncertainty in productivity estimates, and (2) sensitivity
125 of these estimates to uncertainty in each trait. We examine model performance under
126 the estimated uncertainty of each required life history parameter for ten populations of
127 comparatively well-studied ground sharks (order Carcharhiniformes) found in the Gulf
128 of Mexico and northwest Atlantic Ocean. We selected these species as they are well
129 studied and there is relatively good life history information for them. Specifically, we
130 calculate uncertainty in r_{max} estimates through Monte Carlo resampling from probability
131 distributions of the three input parameters required in the model: annual reproductive
132 output, age at maturity, and instantaneous natural mortality. To assess sensitivity, we
133 also compare models that only include uncertainty from individual life history traits.

134 2 Methods

135 [Figure 1 about here.]

136 We used a Monte Carlo simulation model (Fig. 1) based on published information
137 on the biology of a species to iteratively estimate maximum intrinsic rate of population
138 increase r_{max} using a derivation of the Euler-Lotka model (Cortés, 2016; Pardo *et al.*,

139 2016). The model starts with the data required (Values for age at maturity α_{mat} , maxi-
140 mum age α_{max} , litter size l , and breeding interval i (“Data” section in Fig. 1), which are
141 then used to define probability distributions for each parameter (except breeding interval
142 whose value is fixed, see “Probability distributions” in Fig. 1). Values for age at maturity,
143 maximum age, and litter size are then drawn from these distributions (see “Parameters
144 drawn” in Fig. 1), and used to estimate natural mortality M and annual reproductive
145 output b (see “Model inputs” section), which in turn are required to obtain an estimate
146 of r_{max} (blue box in Fig. 1). The drawing of parameters from distributions is repeated
147 20,000 times to obtain 20,000 r_{max} estimates (see inner loop in Fig. 1). Finally, we re-
148 peat the whole process after replacing the probability distributions of each parameter
149 with a fixed value to assess the sensitivity of the model to uncertainty in α_{mat} , b , and
150 M (see outer loop in Fig. 1). We apply this model to ten populations of ground sharks
151 (order Carcharhiniformes) to examine how the uncertainty in traits underlies uncertainty
152 in r_{max} .

153 2.1 Life history data

154 The population-specific life history information required for this simulation model consists
155 of age at maturity (range of years), maximum age (in years), ranges of litter size (in
156 number of female pups), and breeding interval (in years) (Table 1, Fig. 1 “Data” section).

157 [Table 1 about here.]

158 The annual reproductive output of females (b) was calculated as $b = 0.5 * l / i$, where l
159 is litter size (in numbers of males and females) and i is breeding interval (in years). An-
160 nual reproductive output estimates were derived from uniform distributions constrained
161 by the minimum and maximum litter sizes published in the literature. This was neces-
162 sary because empirical distributions of litter sizes are lacking for most elasmobranchs.
163 We assumed a sex ratio of 1:1 to calculate numbers of females only per litter. Age at
164 maturity (α_{mat}) estimates were derived from uniform distributions constrained between
165 the minimum and maximum ages at maturity published in the literature (Table 1).

166 Instantaneous natural mortality M was estimated using the reciprocal of average lifes-
167 pan ($M = 1/\omega$), with average lifespan ω defined as the midpoint between age at maturity
168 and maximum age ($\omega = \frac{\alpha_{mat} + \alpha_{max}}{2}$), for rationale see Pardo *et al.* (2016). Given that
169 we obtained a distribution of age at maturity values for each population (see above),
170 we used this uncertainty in age at maturity as the basis to estimate uncertainty in M ,
171 thus uncertainty of M was iteratively estimated using the same age at maturity dis-
172 tribution described above. Uncertainty in instantaneous natural mortality M had very
173 little influence on r_{max} compared to the effect of uncertainty in age at maturity α_{mat}
174 and annual reproductive output b ; however, this is an artifact of the constrained range

175 of M values produced by our estimation method. When accounting for uncertainty in
 176 natural mortality using the reciprocal of lifespan equation, we only included uncertainty
 177 in age at maturity as it is difficult to set a plausible range for maximum age. These
 178 narrow estimates of M (see Fig. S2d in Supplementary Materials) resulted in uncertainty
 179 in M having a very small effect on r_{max} estimates (see Results). We could increase the
 180 degree of uncertainty in M estimates, yet this increase would be arbitrary. We explored
 181 an alternative scenario where we arbitrarily increase uncertainty in our M estimate and
 182 as expected, its importance on r_{max} estimates increased considerably (see “Alternative
 183 scenario with a more uncertain M ” in Supplementary Materials). The effect of uncer-
 184 tainty in M on r_{max} estimates will depend on the M estimator used and the degree of
 185 uncertainty associated with it (Then *et al.*, 2015). Because of the difficulties of specifying
 186 an adequate level of uncertainty in M as well as how resource intensive it would be to
 187 obtain better estimates of M , we focus our analysis on the effects of including uncertainty
 188 in the other parameters required to estimate r_{max} : α_{mat} and b .

189 2.2 Estimating r_{max}

190 We estimated the maximum intrinsic rate of population increase r_{max} , which in theory,
 191 occurs at low population sizes (i.e., in the absence of density dependence) whereupon it is
 192 equivalent to the fishing mortality required to drive a population to extinction F_{ext} (Dulvy
 193 *et al.*, 2004; Gedamke *et al.*, 2007; Cortés *et al.*, 2014). These two metrics are equivalent
 194 because, in order for a species to become extinct from fishing, fishing mortality needs
 195 to be to equal or exceed the maximum population growth rate ($F_{ext} \geq r_{max}$). Unlike
 196 previous estimates of r_{max} for chondrichthyans (García *et al.*, 2008; Dulvy *et al.*, 2014b),
 197 this equation accounts for juvenile mortality which has been previously overlooked (Pardo
 198 *et al.*, 2016; Cortés, 2016)

$$l_{\alpha_{mat}} b = e^{r_{max}\alpha_{mat}} - e^{-M}(e^{r_{max}})^{\alpha_{mat}-1} \quad (1)$$

199 where $l_{\alpha_{mat}}$ is survival to maturity in the absence of fishing and is calculated as
 200 $l_{\alpha_{mat}} = (e^{-M})^{\alpha_{mat}}$. Because $l_{\alpha_{mat}}$ is derived from M and α_{mat} we did not examine the
 201 effect of uncertainty in $l_{\alpha_{mat}}$ independently. Equation 1 is equivalent to equation 8 in
 202 (Myers and Mertz, 1998) which assumes age at selectivity (i.e., age at which they begin
 203 to be captured) is 1. We used Monte Carlo simulation to propagate uncertainty of input
 204 parameters. We drew parameters from their respective distributions iteratively 20,000
 205 times, and solved for r_{max} (inner loop in Fig. 1)

206 We calculated two metrics that quantify uncertainty in the estimation of r_{max} . We
 207 first calculated the range between the 2.5% and 97.5% quantiles, which encompasses 95%
 208 of the r_{max} values in each distribution and we refer to as the 95% quantile interval. In
 209 order to compare variation for parameters of different magnitudes, we also calculated a

210 coefficient of variation (CV, %) by dividing the mean by the standard deviation of each
211 distribution.

212 **2.3 Model sensitivity**

213 To assess how uncertainty in each parameter affected estimated of r_{max} , we repeated the
214 Monte Carlo simulation for each of the seven possible model combinations with uncer-
215 tainty in: (i) only b , (ii) only α_{mat} , (iii) only M , (iv) $b + \alpha_{mat}$, (v) $b + M$, (vi) $\alpha_{mat} + M$,
216 and (vii) a full model of $b + \alpha_{mat} + M$ (Fig. 1g).

217 In order to visualise the parameter space of r_{max} values created by ranges of α_{mat} , b ,
218 and M , we created two-dimensional contour plots, showing r_{max} estimates along gradients
219 of α_{mat} and b , plotted separately for three levels of M : low $M = 0.05 \text{ year}^{-1}$, medium M
220 $= 0.1 \text{ year}^{-1}$, and high $M = 0.2 \text{ year}^{-1}$. We chose these three values of M as they span
221 the natural mortality values we estimated for the ten shark populations examined.

222 All models were built in R version 3.2.4 (R Core Team, 2016). The r_{max} equation
223 was solved using the `nlminb` optimisation function by minimising the sum of squared
224 differences.

225 **3 Results**

226 **3.1 Estimation of r_{max} and comparison of uncertainty in input 227 parameters**

228 The median r_{max} estimates were robust to uncertainty in all populations examined likely
229 due to the symmetric uncertainty in the underlying parameters. As expected, uncertainty
230 in r_{max} estimates varied considerably among species as a result of uncertainty in the
231 underlying traits (Fig. 2).

232 Estimates of r_{max} are most sensitive to uncertainty in annual reproductive output b
233 (Fig. 2). This is particularly pronounced in the least fecund species (Fig. 3b), that is,
234 those with ranges of annual reproductive output b less than 5 such as the Finetooth Shark
235 *C. isodon* ($b = 2.7$; Fig. 2b) and Blacknose Shark *C. acronotus* ($b = 1.8$; Fig. 2c) as these
236 species had larger differences in coefficients of variation between b and α_{mat} (Table 2).

237 [Figure 2 about here.]

238 [Figure 3 about here.]

239 By focusing on the Atlantic Sharpnose Shark *R. terraenovae* we see that the un-
240 certainty r_{max} values for a model only incorporating uncertainty in b is approximately
241 twice as great as those in the model only incorporating uncertainty in α_{mat} (Fig. 2i).

242 For this species the coefficient of variation (CV) in r_{max} estimates is 32% when only
 243 accounting for uncertainty in b compared to 10% when only accounting for uncertainty
 244 in α_{mat} (Table 2). While the ranges of plausible age at maturity values were low for all
 245 ten populations examined when compared with variation in the other traits (see Fig. 2b
 246 in Supplementary Materials), they still resulted in considerable uncertainty in r_{max} esti-
 247 mates (Fig. 2). Furthermore, CVs of r_{max} estimates were moderately higher in species
 248 with lower estimates of α_{mat} , which are often thought to be relatively resilient to fishing,
 249 than those with higher α_{mat} estimates (Fig. 3a). The CV of r_{max} estimates increased as
 250 average b values of decreased (Fig. 3b). The species in the lower left corner of Fig. 3b
 251 that have low b values but do not have high variability in r_{max} CV are *C. leucas* and
 252 *C. brevipinna*; these species have the highest minimum litter sizes among those with low
 253 b values (see Table 1). Thus, CV is reduced as parameter draws of litter size are very
 254 constrained and never below six, which result in b estimates that are also constrained
 255 (hence low CV in r_{max} estimates) but that never approximate zero.

256 For all ten populations, the full model incorporating all uncertainties ($b + \alpha_{mat} + M$)
 257 had a slightly smaller CV than the $\alpha_{mat} + b$ model due to the close correlation between
 258 α_{mat} and M (Table 2). This same pattern exists when comparing models with just α_{mat}
 259 versus $\alpha_{mat} + M$.

260 Accounting for uncertainty in life history parameters is important: theoretical bio-
 261 logical reference points based on the 2.5% quantile of r_{max} (equivalent to F_{ext}) were on
 262 average 60% lower when all sources of uncertainty were accounted for than when the
 263 deterministic model was used (Fig. 4).

264 [Figure 4 about here.]

265 [Table 2 about here.]

266 3.2 Model sensitivity

267 The interactive effects of annual reproductive output and age at maturity on r_{max} are
 268 nonlinear and vary based on the values of natural mortality ($M = 0.05, 0.1, 0.2 \text{ year}^{-1}$;
 269 Fig. 5). Overall, r_{max} drops steeply at low b values regardless of α_{mat} or M (bottom left
 270 corner of all plots in Fig. 5). At medium to high values of annual reproductive output,
 271 the estimate of r_{max} becomes increasingly sensitive to variation in age at maturity. With
 272 increasing M values, there are increasing combinations of α_{mat} and b values that result
 273 in implausible r_{max} values ($r_{max} \leq 0$; red areas in Fig. 5b & c). This “implausibility”
 274 threshold is particularly apparent when natural mortality M is less than 0.2 year^{-1} , b is
 275 very low ($b < 1$), and α_{mat} is over 5 years (Fig. 5b & c). Nonetheless, it is highly unlikely
 276 that any species have this combination of life histories.

277 [Figure 5 about here.]

4 Discussion

The availability of simple methods for estimating key population parameters has opened the door to comparative risk assessment of a wider range of bycatch species (Stobutzki *et al.*, 2002; Hobday *et al.*, 2011; Dulvy *et al.*, 2014a). One such simple method is the Euler-Lotka approach to estimating the maximum intrinsic rate of population increase using point estimates of three life history traits (García *et al.*, 2008; Hutchings *et al.*, 2012; Dulvy *et al.*, 2014a). However, all life history parameters are estimated with some associated uncertainty. Here, we show that the degree of uncertainty in life history parameters has a considerable effect on the distribution of the resulting range of maximum intrinsic rate of population increase, but little effect on median values. Fully propagating the uncertainty in natural mortality M , age at maturity α_{mat} , and annual reproductive output b , increased the coefficient of variation of r_{max} values by between 11 and 46% (Table 2). These findings have important implications for the use of r_{max} estimates to set fishing limits for sharks and other data-poor species (Fig. 5). Acknowledging the level of uncertainty associated with estimates is crucial when using the precautionary approach, as the degree of risk associated with specific management practices can be estimated (Harwood and Stokes, 2003; Artelle *et al.*, 2013).

As we have shown, r_{max} is particularly sensitive to differences in annual reproductive output, particularly for species with very low annual reproductive output ($b < 5$ females per year; Fig. 3b). That demography is influenced by fecundity of the least fecund species is apparent from some demographic models, but it depends on how reproductive output is parameterised. Our finding is consistent with that of more data-intensive age-structured models. For example, an age-structured model of dogsharks (order Squaliformes) revealed that biological reference points can be strongly influenced by their low fecundity (Forrest and Walters, 2009). In contrast, we caution that another commonly used unstructured model—the rebound potential model (Au and Smith, 1997)—is agnostic to annual reproductive output as the values of b on both sides of the equation cancel out and hence are not considered mathematically in this model (Au *et al.*, 2015). This difference in the implementation of annual reproductive output (b) between two superficially similar unstructured models may help explain differences in species’ sensitivity between these methods. For example, Ward-Paige *et al.* (2013) used the rebound potential model to compare the sensitivity of manta rays (*Mobula spp.*) to that of other elasmobranchs. They found that manta rays were intermediate in sensitivity (r), are more similar to Spinner Shark (*Carcharhinus brevipinna*) or Silky Shark (*C. falciformis*). However, when the very low annual reproductive output of manta rays is accounted for using r_{max} , they were found to have one of the lowest population growth rates r_{max} observed in chondrichthyans (Dulvy *et al.*, 2014a). The annual reproductive output of manta rays is highly uncertain, but with the potential of skipped mating, reproductive output may be as low as one fe-

316 male pup every second or third year (Couturier *et al.*, 2012; Marshall and Bennett, 2010),
317 resulting in r_{max} varying from 0.089 to 0.139 year⁻¹ (Dulvy *et al.*, 2014a). Being aware of
318 major differences in the implementation of annual reproductive output in different models
319 is important when choosing the model best suited to the research question.

320 We show that the highest demographic uncertainty occurred in species with very low
321 annual reproductive outputs—less than five female pups per year. Many elasmobranchs
322 have this reproductive rate (Cortés, 2000). Nonetheless, we reached this conclusion based
323 on assuming a uniform distribution of litter sizes, but they are unlikely to be uniform
324 in the real world. For us to explore the validity of this assumption requires a better
325 understanding of the empirical distribution of litter sizes. By having more accurate
326 distributions of litter size, the uncertainty in annual reproductive output b is likely to
327 decrease, and thus result in more accurate estimates of r_{max} . Hence, we urge biologists to
328 report frequency distributions of individual litter sizes: choosing adequate distributions
329 based on empirical data will result in reduced uncertainty in annual reproductive output.
330 As we have shown in our study, this has considerable downstream effects in the uncertainty
331 of r_{max} estimates.

332 The updated model for estimating r_{max} includes juvenile survival which is derived
333 from adult natural mortality M (Pardo *et al.*, 2016). Yet, because of the known trade-off
334 between offspring size and litter size (Smith and Fretwell, 1974; Hussey *et al.*, 2010), the
335 least fecund species often have the largest offspring. As is typical for marine fishes, such
336 larger offspring will likely have a greater survival probability than the smaller offspring of
337 species with r_{max} among species. An example of this can be illustrated by comparing the
338 Spinner Shark (*C. brevipinna*) with the Scalloped Hammerhead (*S. lewini*). The Spinner
339 Shark litter size ranges between 3–15 individuals born between 60 and 80 cm in length,
340 while the Scalloped Hammerhead has a larger litter size ranging between 13–41 indi-
341 viduals but which are born smaller, between 45 and 50 cm in length (Last and Stevens,
342 2009). Spinner Sharks likely have lower young-of-year survival rates than Scalloped Ham-
343 merheads, resulting in a higher survival to maturity and thus higher r_{max} estimates than
344 those from our study. Additionally, local-scale differences in habitat have also been shown
345 to relate with M (Heupel and Simpfendorfer, 2011). Nuances in relation to differential
346 juvenile mortality among species are not accounted for in our model but should be the
347 focus of further study.

348 Natural mortality is one of the most important parameters in fisheries modelling
349 but one of the hardest to estimate (Pope, 1975; Vetter, 1988; Kenchington, 2014). Our
350 estimates of r_{max} are relatively insensitive to uncertainty in M for shark-like life histories
351 (Figs. 2) because of the method we used for accounting for uncertainty in M . For example,
352 the CV of r_{max} estimates for *R. terraenovae* is 1% when accounting for uncertainty in
353 only M yet it increases to 31.5% if only uncertainty in annual reproductive output is
354 taken into account (Table 2). Our finding is similar to Au *et al.* (2015), who showed that

355 M had only a minor role in the estimation of rebound potential when compared with
356 α_{mat} . However, as natural mortality is a difficult parameter to estimate it is unrealistic
357 for its uncertainty to be narrowly constrained as it was with our method. This begs the
358 question of how to assess how much uncertainty in M is “enough” or “too much”.

359 Furthermore, there is considerable debate as to which empirical model should be used
360 to estimate natural mortality M (Hewitt and Hoenig, 2005; Quiroz *et al.*, 2010; Then
361 *et al.*, 2015). Here, we used the reciprocal of average lifespan; however, other methods,
362 such as Chen and Watanabe (1989) and Peterson and Wroblewski (1984), have been used
363 as they provide varying values of M through ontogeny as required for age-structured
364 demographic modelling (Pardo *et al.*, 2012). We note the differences in r_{max} estimates in
365 our study compared to Cortés (2016) arise because Cortés uses the minimum estimate of
366 a range of mortality estimators, whereas we use one simple method for consistency and
367 transparency.

368 Improving natural mortality estimates, and identifying the best indirect estimator,
369 would require the use of data-intensive methods such as extensive catch data to analyse
370 catch curves, mark recapture experiments, virtual population analyses (VPA), or even
371 fully integrated stock assessments (Kenchington, 2014; Sippel *et al.*, 2017). Therefore,
372 research efforts on data-poor species might be better spent improving on life history
373 estimates of the other parameters age at maturity and annual reproductive output. Fur-
374 ther increasing uncertainty in natural mortality (by an arbitrary amount) does increase
375 uncertainty in r_{max} (see Supplementary Materials). While in our study M is relatively
376 unimportant, more complex age- and stage-structured models consistently show that ju-
377 venile mortality has important contributions to population growth rate (Cortés, 2002;
378 Frisk *et al.*, 2005) and realized recovery potential (Hutchings and Kuparinen, 2017). This
379 difference in importance of M needs to be borne in mind when comparing across demo-
380 graphic model types.

381 While age-dependent mortality estimates are needed for matrix models, only average
382 natural mortality values (and the uncertainty around them) are needed for unstructured
383 models. Natural mortality estimates in which correlation of parameters can be accounted
384 for (e.g., Pauly, 1980) reduced uncertainty of estimates and their error when applied to
385 elasmobranchs (Quiroz *et al.*, 2010). A recent study by Then *et al.* (2015) suggests that
386 the Pauly (1980) mortality model should not be used, and instead a new variant that
387 eliminates temperature from the equation is preferred. Regardless, our results showing
388 the effect of uncertainty in age at maturity and annual reproductive output on r_{max}
389 estimates are likely to be robust to the choice of natural mortality estimator used.

390 We show that accounting for uncertainty in trait estimates is important as it can lead
391 to much more conservative estimates of fishing limits than if uncertainty is ignored. In-
392 corporating uncertainty also considerably increased the potential range of maximum popu-
393 lation growth rate r_{max} estimates in these relatively well-studied sharks. Furthermore, we

394 urge managers and policy makers to focus on restricting mortality in species with annual
395 reproductive outputs <2 females per year, particularly if this mortality occurs among
396 younger age classes. For data-limited species, we recommend incorporating all sources
397 of uncertainty in life history parameters, especially when lacking directly-estimated life
398 history parameters. This uncertainty in r_{max} can be reduced by understanding the corre-
399 lation in life history parameters, and we encourage researchers studying chondrichthyan
400 reproductive biology to report distributions of litter sizes and offspring sizes as to reduce
401 uncertainty in these parameter distributions and thus better quantify known uncertainty
402 in demographic productivity models.

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Table 1: Values and sources of life history parameters used to estimate r_{max} for the ten populations studied. Note that annual reproductive output b is not obtained directly from the literature but is estimated from litter size and breeding interval.

Species	Litter size		Breeding interval	Mean b	Age at mat.		Max. age	Source
	min	max			min	max		
<i>Carcharhinus acronotus</i>	1	6	2.0	1.8	4.0	5.0	19.2	Driggers <i>et al.</i> (2004a,b); Barreto <i>et al.</i> (2011); Branstetter (1990)
<i>Carcharhinus brevipinna</i>	6	10	2.0	4.0	7.0	8.0	16.0	Cortés (2002); Branstetter (1987a)
<i>Carcharhinus isodon</i>	2	6	1.5	2.7	3.3	5.3	8.0	Castro (1993); Carlson <i>et al.</i> (2003); Driggers and Hoffmayer (2009)
<i>Carcharhinus leucas</i>	6	12	2.0	4.5	17.0	19.0	31.0	Branstetter (1990); Cliff and Dudley (1991); Cortés (2002); Branstetter and Stiles (1987)
<i>Rhizoprionodon terraenovae</i>	1	12	1.0	6.5	2.8	3.9	9.0	Parsons (1983); Bigelow and Schroeder (1948); Branstetter (1987b); Parsons (1985)
<i>Sphyrna lewini</i>	12	38	1.0	25.0	13.0	15.0	36.0	Branstetter (1987c); Drew <i>et al.</i> (2015); Stevens and Lyle (1989); Cortés (2002)
<i>Sphyrna mokarran</i>	13	42	2.0	13.8	7.4	9.5	31.7	Harry <i>et al.</i> (2011); Compagno (1984); Stevens and Lyle (1989); Last and Stevens (2009)
<i>Sphyrna tiburo</i>	3	15	1.0	9.0	2.9	4.0	7.5	Lombardi-Carlson <i>et al.</i> (2003); Cortés (2002)
<i>Carcharhinus limbatus</i> ATL	2	10	2.0	3.0	5.7	7.7	21.6	Carlson <i>et al.</i> (2006); Branstetter (1990); Castro (1996)
<i>Carcharhinus limbatus</i> GULF	2	10	2.0	3.0	4.7	6.7	14.4	Carlson <i>et al.</i> (2006); Branstetter (1990); Castro (1996)

Table 2: Coefficients of variation (CV) in r_{max} estimates of seven models with varying levels of uncertainty in age at maturity, natural mortality and annual reproductive output for the ten shark populations examined. Means, medians and standard deviations are also presented for the full model which accounts for uncertainty in all three parameters.

Species	Model Coefficient of Variation (CV, %)										$b + \alpha_{mat} + M$ (Full) Model	
	b	α_{mat}	M	$b + \alpha_{mat}$	$b + M$	$\alpha_{mat} + M$	$b + \alpha_{mat} + M$	Mean	Median	St.Dev		
<i>Carcharhinus isodon</i>	25.0	16.4	4.1	30.0	25.4	12.3	27.9	0.197	0.188	0.061		
<i>Carcharhinus limbatus</i> GULF	25.2	10.2	1.4	27.3	25.2	8.8	26.8	0.195	0.195	0.015		
<i>Rhizoprionodon terraenovae</i>	31.5	9.6	1.0	33.1	31.6	8.7	32.8	0.185	0.185	0.052		
<i>Carcharhinus limbatus</i> ATL	21.8	7.9	0.7	23.2	21.8	7.2	23.0	0.112	0.112	0.009		
<i>Carcharhinus acronotus</i>	31.9	5.5	0.5	32.5	31.9	4.9	32.4	0.428	0.405	0.133		
<i>Carcharhinus brevipinna</i>	7.1	3.9	0.5	8.1	7.1	3.4	7.8	0.240	0.239	0.021		
<i>Carcharhinus leucas</i>	7.4	3.3	0.4	8.1	7.4	2.8	7.9	0.316	0.315	0.037		
<i>Sphyrna tiburo</i>	21.0	9.8	1.1	23.2	21.0	8.7	22.8	0.462	0.453	0.103		
<i>Sphyrna mokarran</i>	9.7	6.6	0.2	11.7	9.7	6.4	11.6	0.202	0.197	0.045		
<i>Sphyrna lewini</i>	7.9	4.0	0.2	8.8	7.9	3.8	8.7	0.204	0.199	0.053		

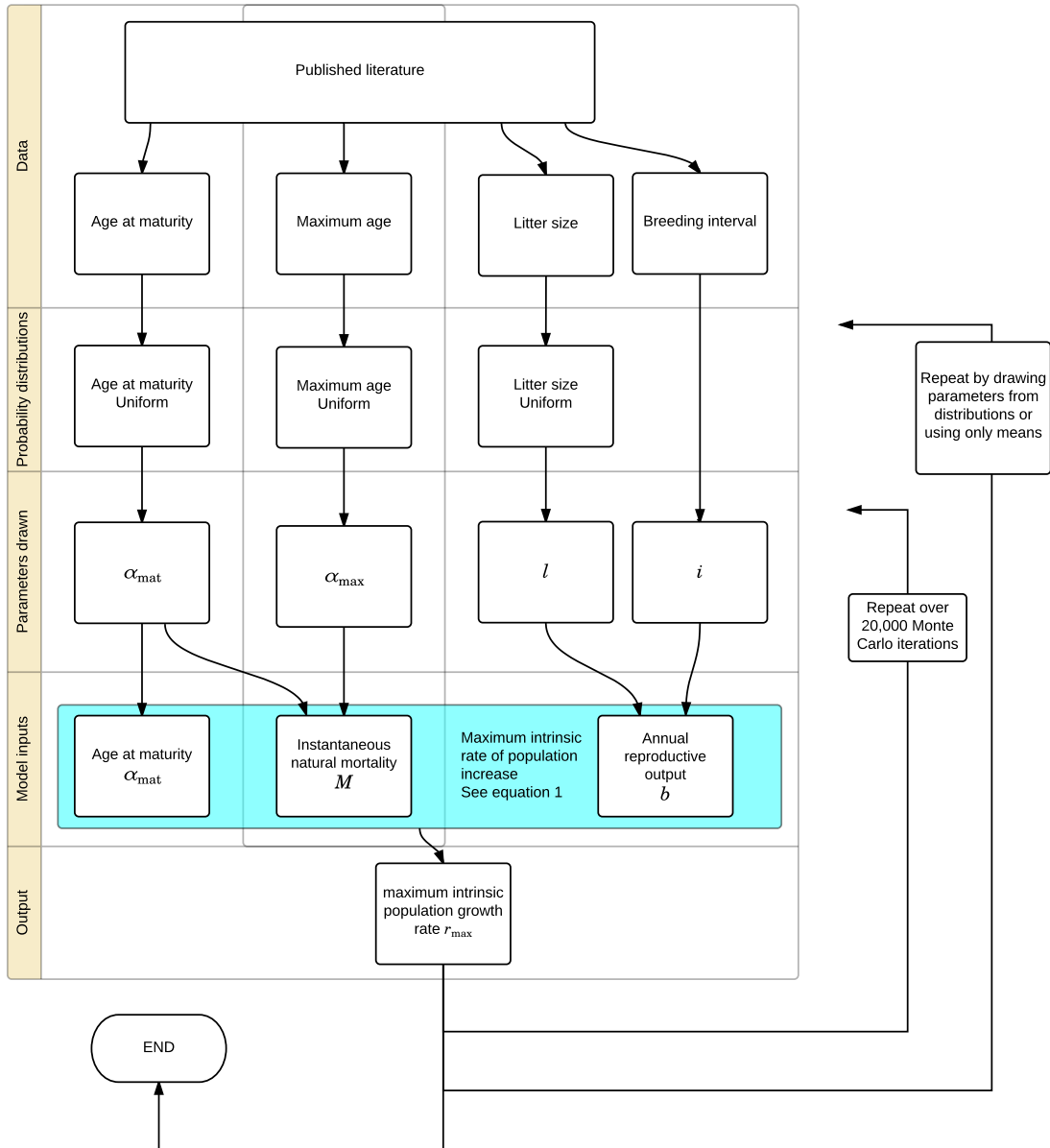


Figure 1: Flow chart illustrating the structure of the Monte Carlo simulation model used in this study. The model starts with the data required (Values for age at maturity α_{mat} , maximum age α_{max} , litter size l , and breeding interval i), which are then used to define probability distributions for each parameter (except breeding interval whose value is fixed). Values for age at maturity, maximum age, and litter size are then drawn from these distributions, and used to estimate natural mortality M and annual reproductive output b , which in turn are required to obtain an estimate of maximum intrinsic rate of population increase r_{max} . The drawing of parameters from distributions is repeated 20,000 times to obtain 20,000 r_{max} estimates. Finally, we replace the probability distributions of each parameter with a fixed value to assess the sensitivity of the model to uncertainty.

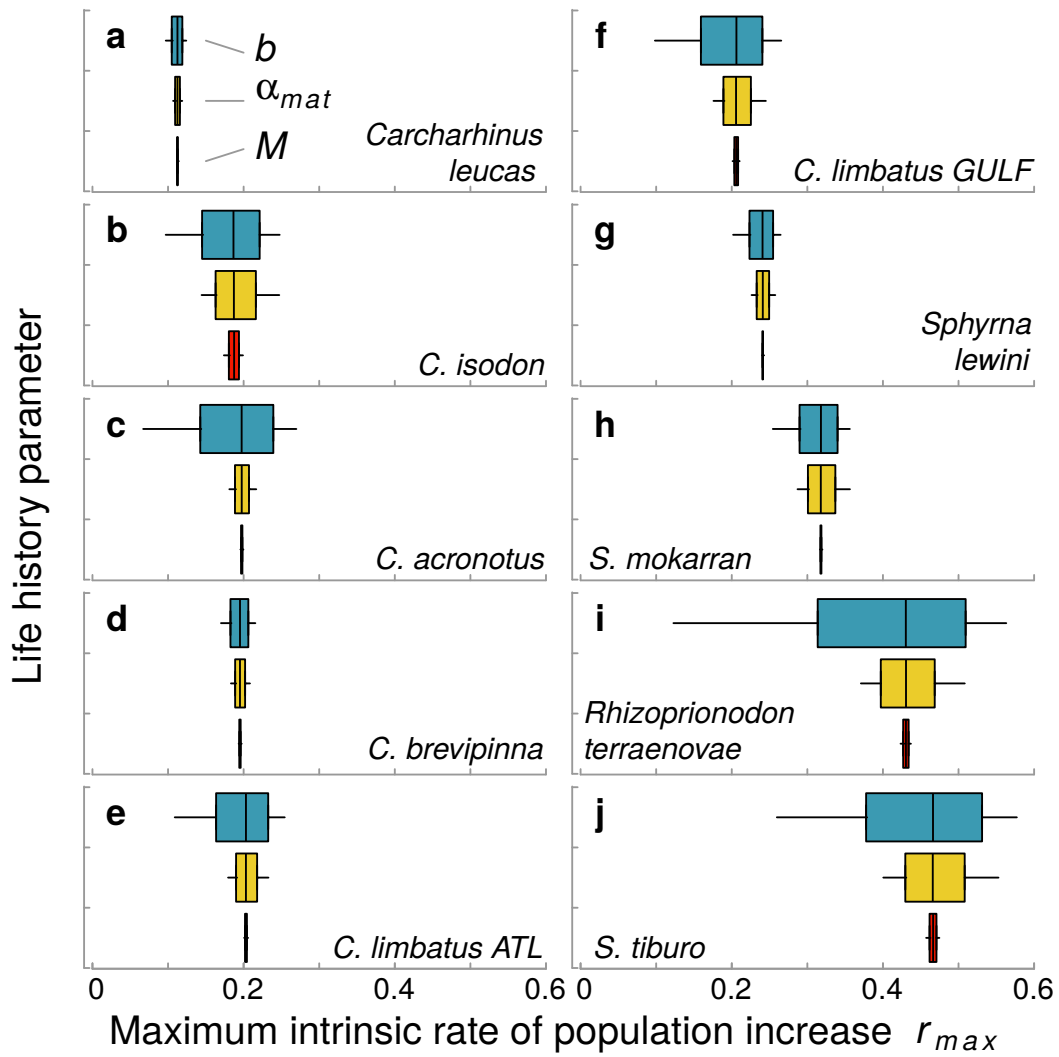


Figure 2: Predicted values of maximum intrinsic rate of increase r_{max} for ten different shark populations when including uncertainty in annual reproductive output b (blue box plots), age at maturity α_{mat} (yellow box plots), and natural mortality M (red box plots). Boxes indicate median, 25% and 75% quantiles, while the lines encompass 95% of the values (2.5% and 97.5% quantiles).

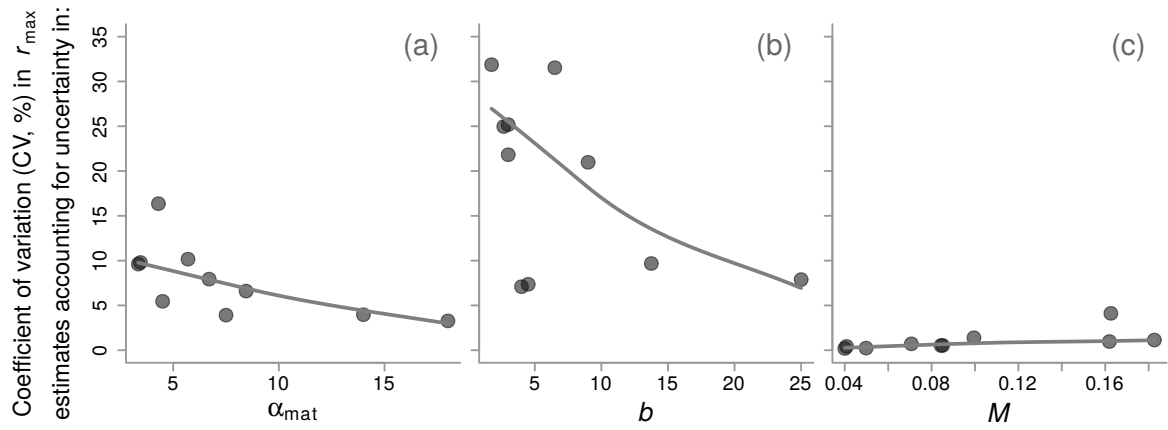


Figure 3: Coefficient of variation (CV, %) in r_{max} estimates for ten different shark populations when accounting for uncertainty in (a) age at maturity α_{mat} , (b) annual reproductive output b , and (c) natural mortality M , plotted against the median values of the respective life history parameter. Lines are loess-smoothed curves.

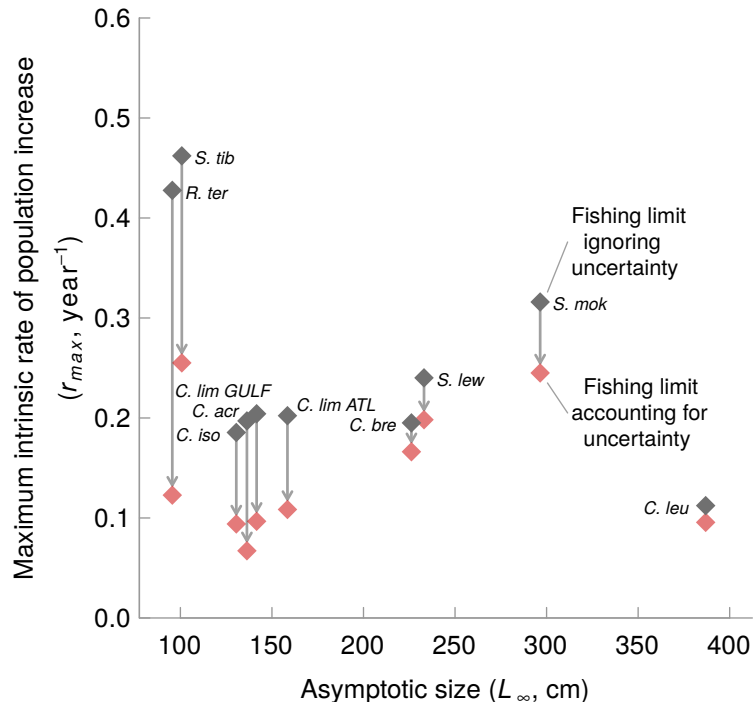


Figure 4: Comparison of potential fishing limits based on r_{max} when estimated with no uncertainty accounted for (grey diamonds) and when accounting for uncertainty from all sources in the model (using 2.5% quantile, red diamonds). Points were slightly jittered horizontally to avoid overlap. Note that r_{max} is equivalent to F_{ext} , i.e. the point at which fishing mortality is equal to the maximum rate of population increase.

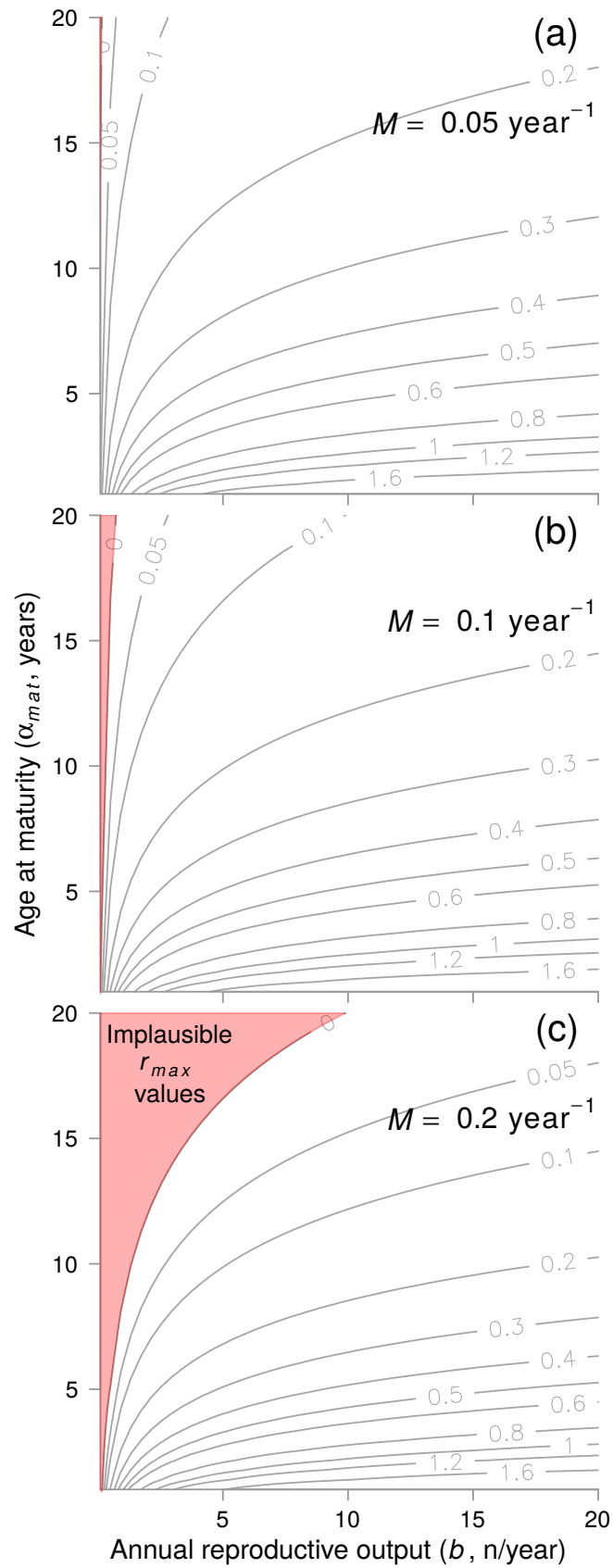


Figure 5: Contour plots of r_{max} values for varying ranges of age at maturity α_{mat} and annual reproductive output b , with values of instantaneous natural mortality M set as (a) low (0.05 year^{-1}), (b) medium (0.1 year^{-1}), and (c) high (0.2 year^{-1}).